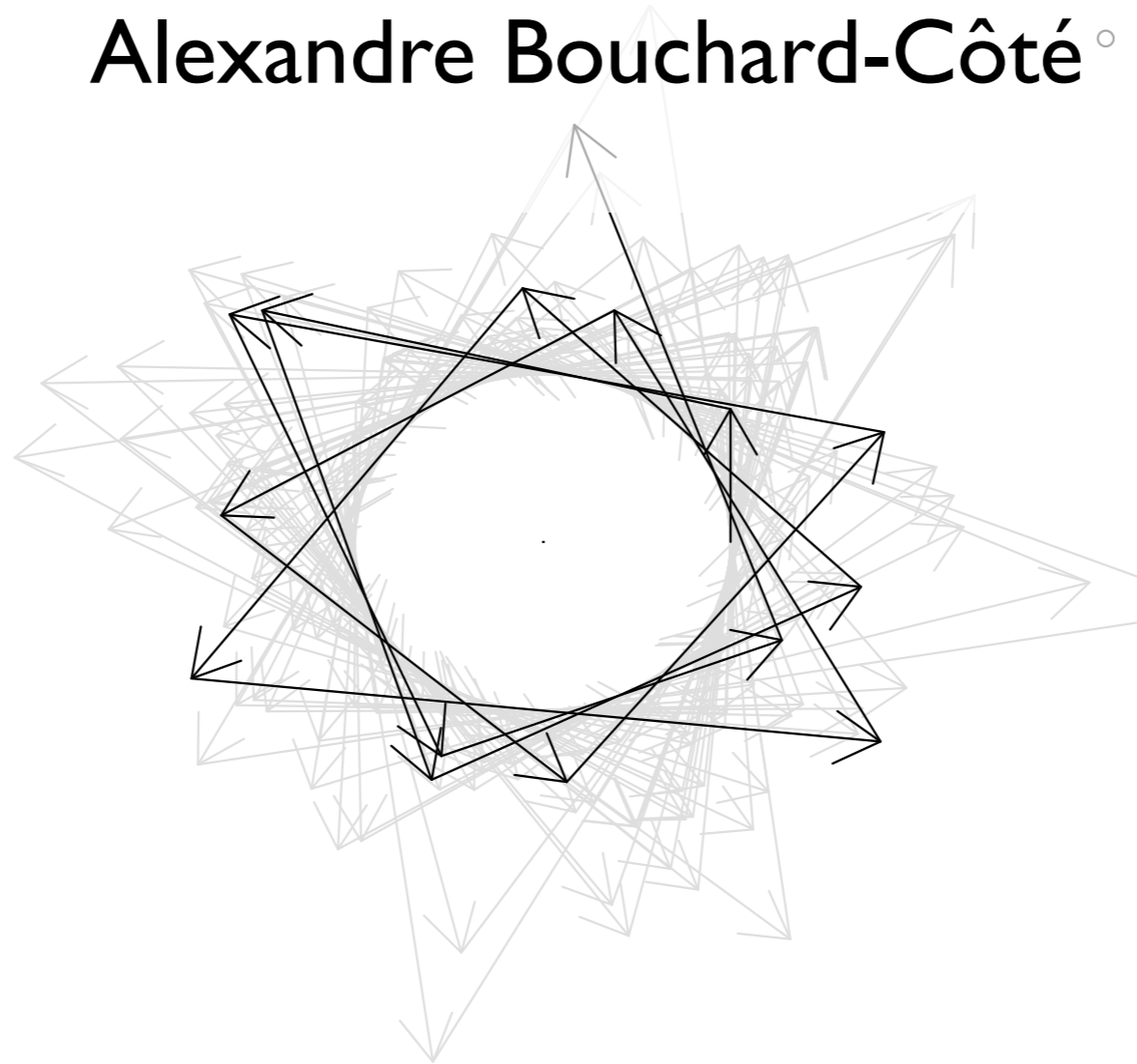
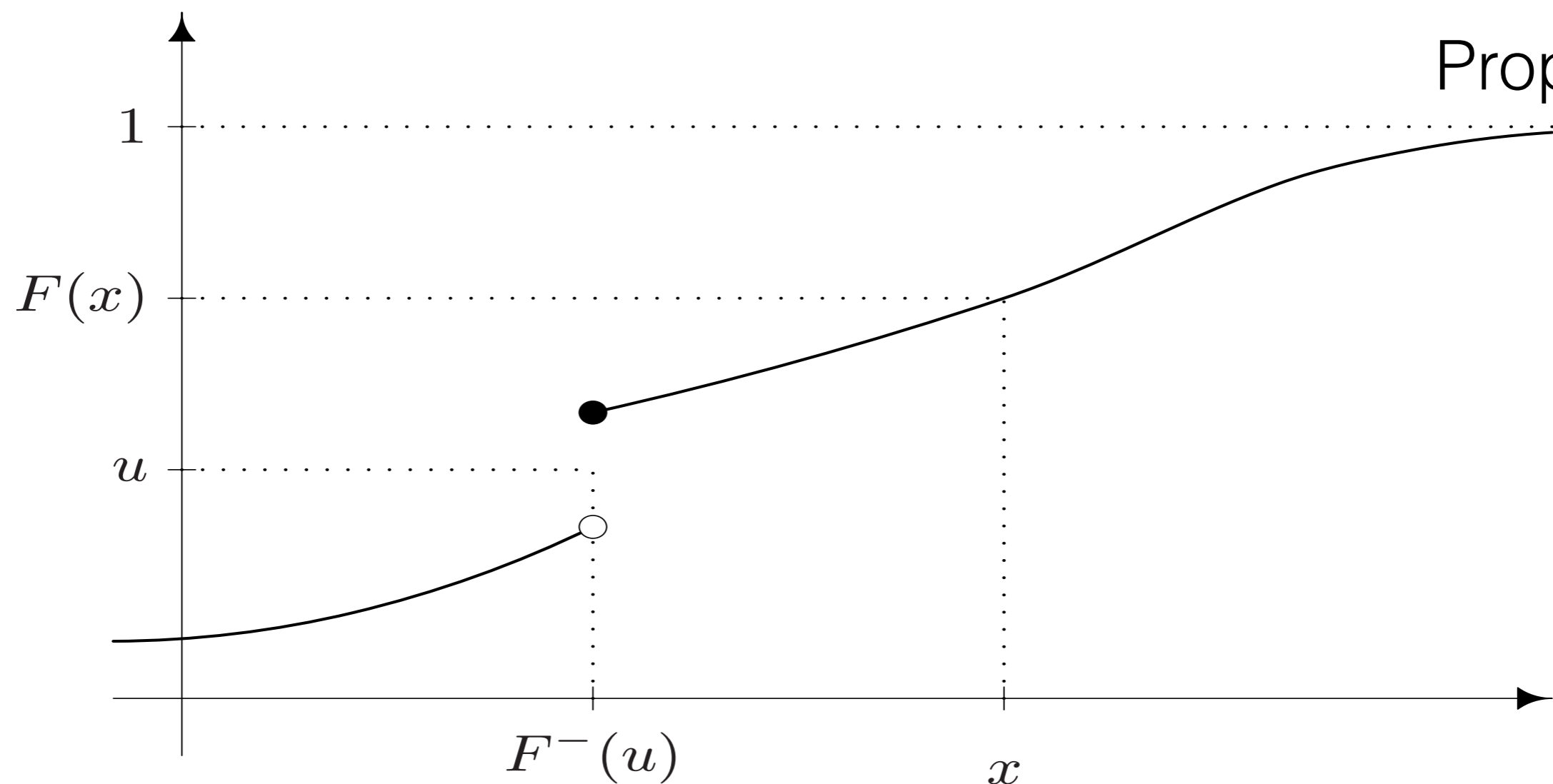


Monte Carlo methods

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**Review / Q&A /
Exercise solutions**



- **Proposition.** Let F be a cdf and $U \sim \mathcal{U}_{[0,1]}$. Then $X = F^{-1}(U)$ has cdf F .

Proof. $F^{-1}(u) \leq x \Leftrightarrow u \leq F(x)$ so $U \sim \mathcal{U}_{[0,1]}$, we have

$$\mathbb{P}(F^{-1}(U) \leq x) = \mathbb{P}(U \leq F(x)) = F(x).$$

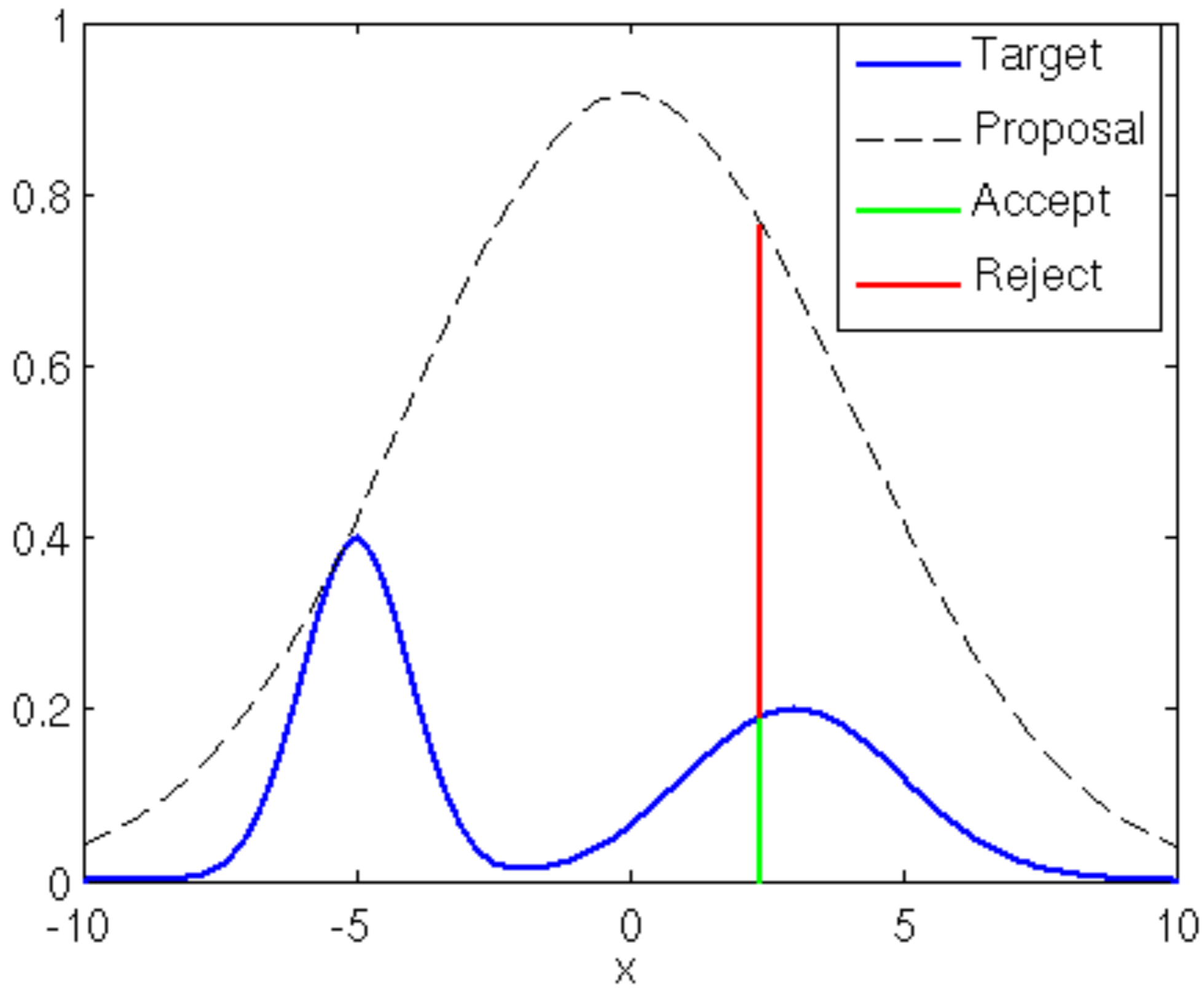
Exercise: construct a RNG for exponential random variables of a given rate

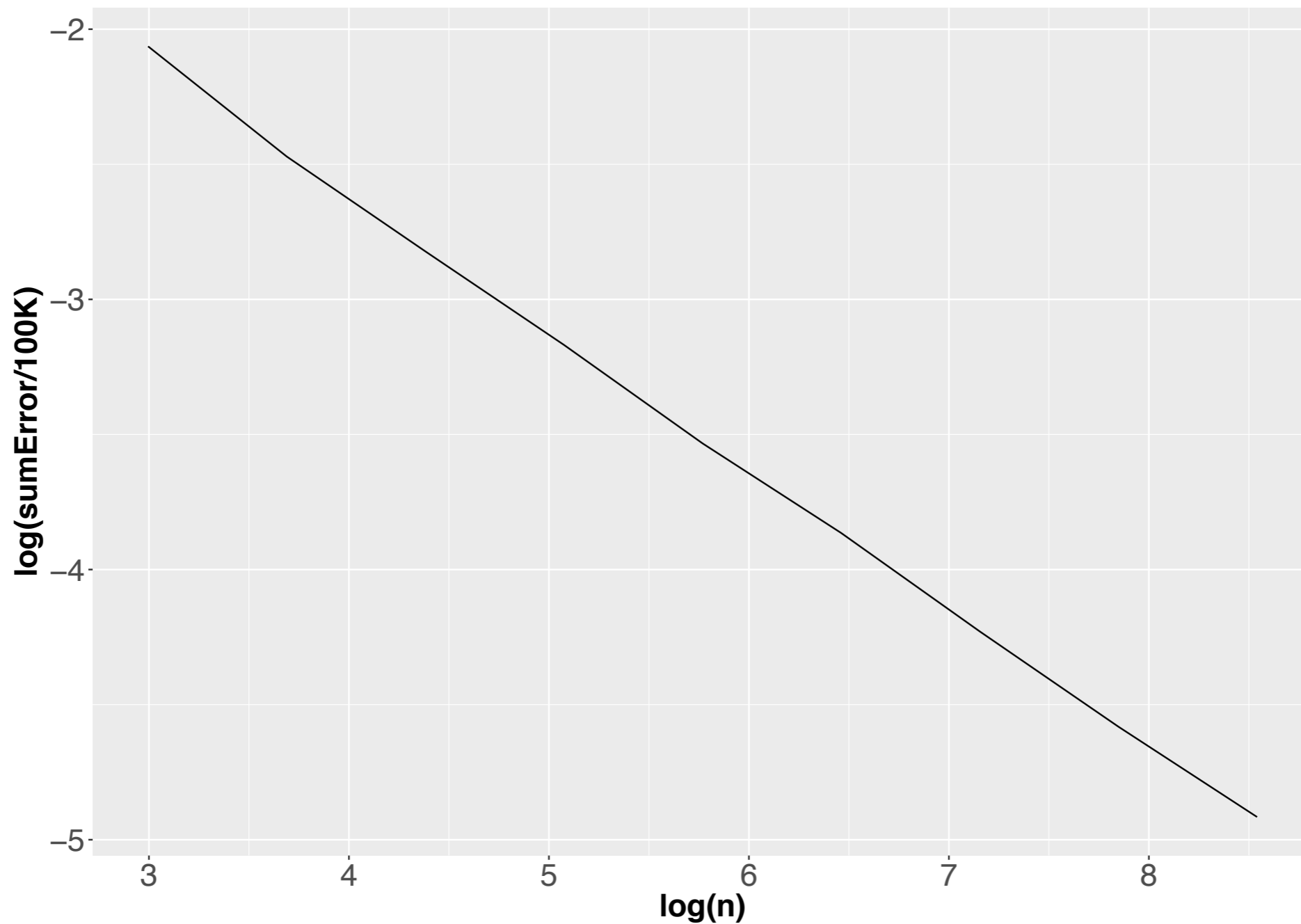
Transformation method

- Let $Y \sim q$ be a \mathbb{Y} -valued random variable (rv) of, which we can simulate (eg, by inversion)
- Let $X \sim \pi$ be a \mathbb{X} -valued rv, which we wish to simulate.
- It may be that we can find a function $\varphi : \mathbb{Y} \rightarrow \mathbb{X}$ with the property that if we simulate $Y \sim q$ and then set $X = \varphi(Y)$ then we get $X \sim \pi$.
- Inversion is a special case of this idea.

Augmentation / auxiliary variables Def. 16

- Example: how to simulate from a mixture model?
- **Key idea:** marginalization is easy with Monte Carlo methods
- Contrast with analytical marginalization
- Build (X, Y) such that distribution of interest is a marginal (or a conditional on a discrete variable)
- **Exercises:**
 - write pseudo-code to simulate from a mixture distribution with 2 normal components,
 - show that rejection sampling is an augmentation sampling scheme





1. Use this plot to empirically derive the running time of Simple Monte Carlo for a given tolerance tol . Create the plot for Example 7 [c.f. lecture slides].
4. Compute the running time in tol and d for Example 7 [c.f. lecture slides] but with non-diagonal covariance normal vectors.

Fundamental equation to analyze Monte Carlo methods

$$\text{running time} = \boxed{\begin{array}{l} \text{number of samples} \\ \text{needed to get a} \\ \text{tolerance (with} \\ \text{probability 95\%)} \end{array}} \times \boxed{\begin{array}{l} \text{compute cost per} \\ \text{sample} \end{array}}$$

- Exercise:
- Compute the running time in tol and d for Example 7 but with non-diagonal covariance normal vectors

Using CLT in practice

- Let X be a random tree
- We are looking at a clade indicator $f(X)$ as in Example 2.
- After 500 iid trees your MC estimate for the clade support is roughly 10%
- Should you extract more samples?
- Say we want a scheme with relative error of less than 10% for approximately 95% of the random seeds

Importance sampling (IS)

IS: when proposal and target have known normalizations

If: $\pi(x) > 0 \Rightarrow q(x) > 0$

Then:

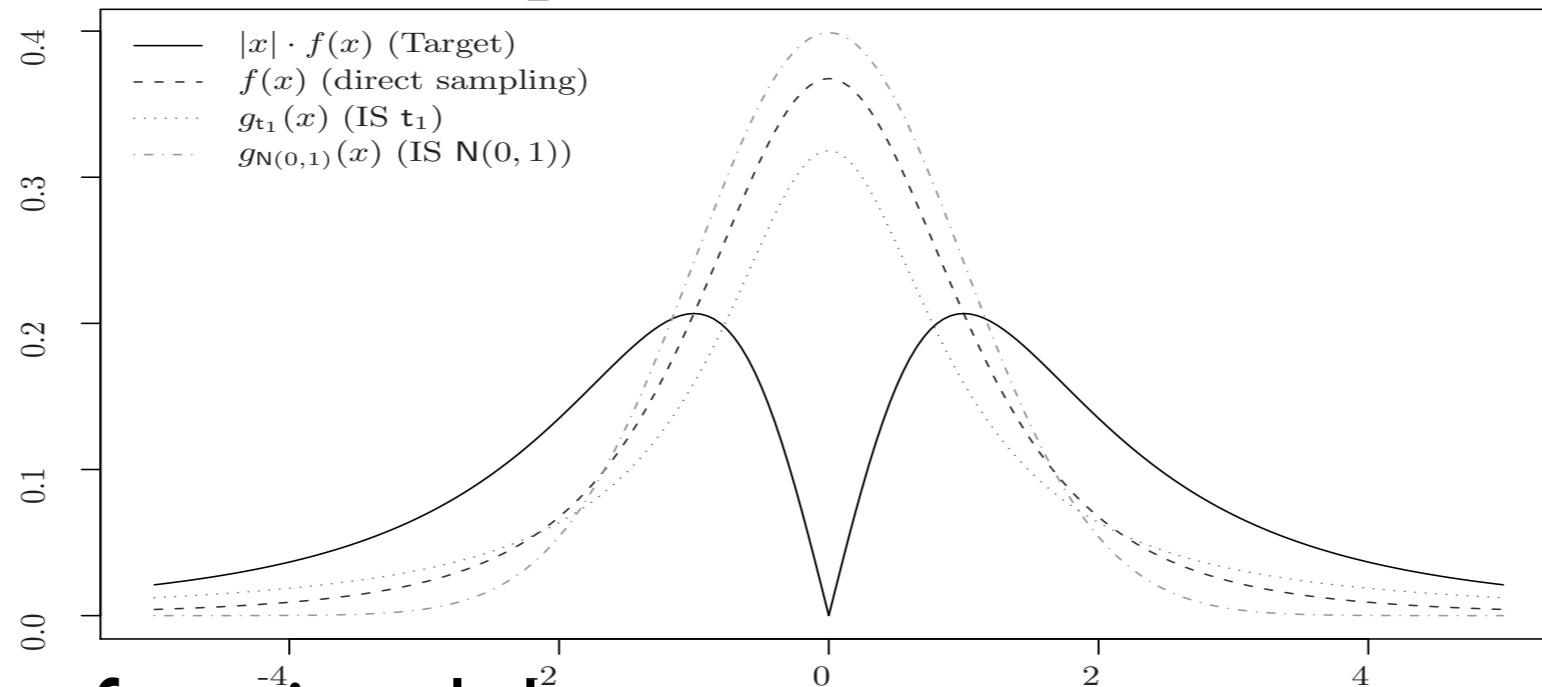
$$\begin{aligned}\pi(A) &:= \int_A \pi(x) dx \\ &= \int_A \underbrace{\frac{\pi(x)}{q(x)}}_{:=w(x)} q(x) dx \\ &= \int_A w(x) q(x) dx\end{aligned}$$

X_1, \dots, X_n be a sample of independent random variables distributed according to q

$$\widehat{I}_n^{\text{IS}} = \frac{1}{n} \sum_{i=1}^n \phi(X_i) w(X_i)$$

Example/exercise

Ex 19



- Test function: $|x|^2$
- Target density: t-distribution, 3 degrees of freedom
- Compare (x-axis, -1 - 1500 , y axis, partial sum, range of 100 replicates)
 - Simple MC
 - IS with t proposal, 1 degree of freedom
 - IS with normal proposal

Normalization: important issue in large scale Bayesian or random effects models

- We are given a density known up to a normalization constant

$$\pi(x) = \frac{\gamma(x)}{Z}$$

- Example: $\pi(x) = \frac{\text{joint}(x, y)}{\text{evidence}(y)}$

x : unknown
 y : data

- We want a law of large number

$$\frac{1}{N} \sum_{i=1}^N \varphi(X^{(i)}) \rightarrow \int \varphi(x) \pi(dx) \text{ a.s.}$$

← *'test function'*

Note: We almost never care about the samples themselves!

IS: when proposal and target *do not* have known normalizations

If: $\pi(x) > 0 \Rightarrow q(x) > 0$

Then:

$$\begin{aligned}
 I &= \mathbb{E}_\pi(\phi(X)) = \int_{\mathbb{X}} \phi(x) \pi(x) dx \\
 &= \frac{\int_{\mathbb{X}} \phi(x) w(x) q(x) dx}{\int_{\mathbb{X}} w(x) q(x) dx} \\
 &= \frac{\mathbb{E}_q(\phi(X)w(X))}{\mathbb{E}_q(w(X))}.
 \end{aligned}$$

X_1, \dots, X_n be a sample of independent random variables distributed according to q

$$\widehat{I}_n^{\text{NIS}} = \frac{\sum_{i=1}^n \phi(X_i)w(X_i)}{\sum_{i=1}^n w(X_i)}$$

Analysis of the asymptotic variance

Assume that $\mathbb{V}_q(\phi(X)w(X)) < \infty$ and $\mathbb{V}_q(w(X)) < \infty$ then

$$\sqrt{n} \left(\hat{I}_n^{\text{NIS}} - I \right) \xrightarrow{D} \mathcal{N} \left(0, \sigma_{\text{NIS}}^2 \right)$$

Exercise: compute asymptotic variance

Tool: delta method

If: $\sqrt{n} (Z_n - \mu) \xrightarrow{D} \mathcal{N} (0, \Sigma) .$

Then: $\sqrt{n} (g(Z_n) - g(\mu)) \rightarrow \mathcal{N} (0, \nabla^T g(\mu) \Sigma \nabla g(\mu)) .$