

# STAT 530: Decision Theory Continued

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# Start with loss function, $L(\theta, a)$ , recall definitions:

- $R(\theta, \delta)$  is **risk function** of rule (estimator)  $\delta()$ :
  - expected value of  $L(\theta, \delta(Y))$ , w.r.t. sampling ( $Y|\theta$ )
- $r(\pi, \delta)$  is **Bayes risk** of  $\delta()$ :
  - average risk w.r.t.  $\theta \sim \pi$
  - think of  $\pi$  as *nature's prior*.
- $\rho_\pi(a; y)$  is **posterior risk** of action (estimate)  $a$ :
  - average of  $L(\theta, a)$  w.r.t. posterior on ( $\theta|Y = y$ )
  - here think of  $\pi$  as *investigator's prior*

# Big Fact #1: Choice of loss function influences choice of which Bayesian estimator

General Bayesian principle: 'the' Bayesian estimator is:

$$\delta_B(y) \equiv \operatorname{argmin}_a E\{L(\theta, a) | Y = y\}$$

Examples:

- $L(\theta, a) = (a - \theta)^2$  implies  $\delta_B(y) = E(\theta | Y = y)$
- $L(\theta, a) = |a - \theta|$  implies  $\delta_B(y) = \operatorname{Median}(\theta | Y = y)$

## Big Fact #2: Sense in which Bayes estimators are best possible

The Bayes risk based on nature's prior  $\pi_N$  is minimized - amongst all possible estimation procedures  $\delta()$  - by the Bayes estimator which uses  $\pi_N$  as the investigator's prior.

Proof by changing the order of integration.

# Big Fact # 3: There is no point in considering non-Bayesian estimators

Recall:  $\delta()$  **inadmissible** if there exists  $\delta^*()$  such that

$$R(\theta, \delta) \leq R(\theta, \delta^*)$$

for all  $\theta$ , with strict inequality for at least one  $\theta$ .

And an estimator is **admissible** if it is not inadmissible.

It is *almost* the case that:

$\delta()$  is admissible  $\leftrightarrow$   $\delta()$  is a Bayes estimator w.r.t. some prior.

Say  $\delta()$  is Bayes estimator w.r.t.  $\pi$ , show  $\delta()$  admissible

Look at simple case  $\theta \in \{\theta_1, \theta_2, \dots\}$

Say  $\delta()$  is admissible, show  $\delta()$  is Bayes w.r.t. some prior  $\pi$

Proof by picture in very simple case  $\theta \in \{\theta_1, \theta_2\}$ :

Visualize any  $\delta()$  as point  $(R_1, R_2) = R(\theta_1, \delta), R(\theta_2, \delta)$  in 'risk set'.

- risk set necessarily convex (think of randomized rules)
- 'lower boundary' of the set corresponds to admissible rules
- at any boundary point, convex set 'supported' by a line
- express line as  $\pi_1 R_1 + \pi_2 R_2 = c$
- Voila! rule minimizes Bayes risk w.r.t. this  $\pi$

# More precisely, admissible estimators are Bayes estimators plus some 'limits' of Bayes estimators

E.g.  $Y \sim N(\theta, 1)$ ,  $L(\theta, a) = (a - \theta)^2$ .

$\delta(Y) = Y$  is

- admissible
- not the posterior mean arising from any proper prior
- the  $\tau^2 \rightarrow \infty$  limit of  $E(\theta|Y = y)$  via the prior  $\theta \sim N(0, \tau^2)$ .



## But things get crazy...

$$Y \sim N_p(\theta, I_p), L(\theta, a) = \|a - \theta\|^2$$

Consider  $\delta(Y) = Y$

- $p = 1$ : admissible
- $p = 2$ : admissible
- $p \geq 3$ : **inadmissible**
  - James-Stein (1960)
  - shocking!

Showed that  $\delta_{JS}(Y) = (1 - (p - 2)/\|Y\|^2)Y$  beats  $\delta(Y) = Y$

Later, someone showed  $\delta_{JS}()$  is beaten by  
 $\delta_{JS}^+(Y) = \max\{0, 1 - (p - 2)/\|Y\|^2\}Y$ .

Later someone showed  $\delta_{JS}^+$  inadmissible,  
*without actually finding an estimator that beats it!*

Isn't math great?