

STATISTICS 536B, Lecture #4

March 5, 2015

Recap: Interested in population-level association between (X, Y) .

Various study designs:

- prospective, sample $(Y|X)$
- retrospective, sample $(X|Y)$
- cross-sectional, sample (X, Y)

Same analysis:

| | X=1 | X=0 |
|-----|-----|-----|
| Y=1 | q | r |
| Y=0 | s | t |

$$\begin{aligned}\widehat{\log OR} &= \log q + \log t - \log r - \log s \\ SE \left[\widehat{\log OR} \right] &= \sqrt{1/q + 1/r + 1/s + 1/t}\end{aligned}$$

And with confounders

Logistic regression of Y on $(1, X, C)$ [or more generally basis functions of (X, C)] is appropriate, whether the actual data acquisition is prospective [sampling $(Y | X, C)$], cross-sectional [sampling (Y, X, C)], or retrospective [sampling $(X, C | Y)$]

Caveat about estimating intercept - intuitively sensible - case-control data cannot tell you how common the disease is in the population.

Matched Case-Control Data with binary exposure: Cross-Classify the n **PAIRS** of subjects

| | | Case | | |
|----------|-------------|-------------|---------|---|
| | | Not exposed | Exposed | |
| Control: | Not exposed | a | b | |
| | Exposed | c | d | |
| ----- | | | | n |

Think about the source population having a distribution over disease status, exposure, and matching factors (confounders) jointly: $f(y, x, m)$

Think about sampling a pair of individuals yielding:

$\{(Y_0, X_0, M_0), (Y_1, X_1, M_1)\}$

but with the **constraints** $Y_0 + Y_1 = 1$ and $M_0 = M_1$.

A likelihood for parameters describing the source population based on how the data were actually sampled??? **Activity.**

So we have a likelihood function for β ,
the $(Y, X|M)$ log-OR

$$L(\beta) = (1/2)^a (1/2)^d \{1/(1 + \exp(-\beta))\}^b \{1/(1 + \exp(\beta))\}^c,$$

$$l(\beta) = -b \log(1 + \exp(-\beta)) - c \log(1 + \exp(\beta)) + \text{constant}$$

And we can apply our usual tools to the log-likelihood, to get inference procedures

Also a **score test** - mentioned in reading

General idea of score test. Have log-likelihood $l(\beta)$ for scalar parameter β .

Under the hypothesis $\beta = 0$,

$$\frac{l'(0)}{\sqrt{-l''(0)}} \underset{\sim}{\text{approx}} N(0, 1)$$

Matching is inefficient/efficient when the matching factor(s) is a weak/strong confounder???

```
### set up a population
t <- 100000
set.seed(17)

m.pop <- sample(0:6, size=t, prob=rep(1/7, 7), replace=T)

x.pop <- rbinom(t, size=1,
               prob=expit(logit(.1)+(m.pop/6)*(logit(.9)-logit(.1))))

y.pop <- rbinom(t, size=1, prob=expit(-2 + 0.5*m.pop + 0.5*x.pop))
```

Sample some cases and controls

```
n <- 400
cs <- sample((1:t)[y.pop==1], size=n)

### one possible way to complete the study - unmatched controls
cn.unmt <- sample((1:t)[y.pop==0], size=n)

### another possible way to complete the study - matched controls
cn.mtch <- rep(NA,n)
for (i in 1:n) {
  cn.mtch[i] <- sample((1:t)[(y.pop==0)&(m.pop==m.pop[cs[i]])], size=1)
}
```

Do the matched study analysis

```
dat.pair <- table(x.pop[cn.mtch], x.pop[cs])
```

```
> dat.pair
```

| | 0 | 1 |
|---|----|-----|
| 0 | 66 | 72 |
| 1 | 65 | 197 |

```
> c(log(dat.pair[1,2])-log(dat.pair[2,1]),  
    sqrt(1/dat.pair[1,2] + 1/dat.pair[2,1]))
```

```
0.102 0.171
```

Do the unmatched study analysis

```
y <- y.pop[c(cs, cn.unmt)]
x <- x.pop[c(cs, cn.unmt)]
m <- m.pop[c(cs, cn.unmt)]

ft <- glm(y~x+as.factor(m), family=binomial)

> c(coef(ft)[2], sqrt(vcov(ft)[2,2]))
0.260 0.191
```

And repeat...

