

## Nonlinear Regression Models

$$E(Y|X) = g(X; \theta) + \epsilon,$$

$g$  NOT linear in  $\theta$ .

Fundamentally different (harder) than GLM.

Secs 8.1-8.6 in text.

**Weight loss example:**

$X$  is days since start of weight reduction program

$Y$  is weight on day  $X$

Exponential decay model:

$$Y = \beta_0 + \beta_1 2^{-X/\lambda} + \epsilon,$$

all parameters have clear interpretations.

**“Stormer Viscometer”** (measures viscosity of a fluid) ex.:

$Y$  is time taken for “inner cylinder to perform a fixed number of revolutions in response to an actuating weight.”

$X_1$  is the viscosity

$X_2$  is the actuating weight

$$Y = \frac{\beta_1 X_1}{X_2 - \beta_2} + \epsilon$$

**Calibration idea:** run the experiment for a number of fluids of known viscosity, using a variety of weights. Use these data to fit the model and “learn” the parameters.

Then set to measure viscosities of other fluids, i.e., predict  $X_1$  given  $Y$  and  $X_2$  (interesting issues lurking here...).

**Least-squares fitting** (MLE assuming iid normal error terms)

Write as  $Y_i = g^{(i)}(\theta) + \epsilon_i$

So

$$\hat{\theta} = \operatorname{argmin}_{\theta} \sum_{i=1}^n \left\{ y_i - g^{(i)}(\theta) \right\}^2 .$$

No closed-form solution.

Newton-Raphson: fine (best) if happy to evaluate 1st and 2nd derivatives of  $g^{(i)}(\theta)$ .

Approach based on 1st derivative only?

Notation:

$$g_j^{(i)}(\theta) \equiv \partial g^{(i)}(\theta) / \partial \theta_j$$

Iterative scheme: let  $\hat{\theta}^{(m)}$  be estimate after  $m$  iterations.

One-term Taylor-series approx. to linearize regression function:

$$\hat{\theta}^{(m+1)} = \operatorname{argmin}_{\theta} \sum_{i=1}^n \left\{ y_i - g^{(i)}(\hat{\theta}^{(m)}) - \sum_{j=1}^p g_j^{(i)}(\hat{\theta}^{(m)}) (\theta_j - \hat{\theta}_j^{(m)}) \right\}$$

i.e., get  $\hat{\theta}^{(m+1)}$  as a linear least-squares fit, with a pseudo response vector and a pseudo design matrix.

**General strategy for statistical computing:** replace a hard problem with an iterative sequence of easy problems.

## Convergence issues:

Need to provide  $\hat{\theta}^{(0)}$  before we can start iterating.

As with virtually all optimization problems, a ‘bad’ choice may produce non-convergence.

Possibilities for ‘good’ choices?

Scientific knowledge?

Somehow use the data to get a reasonable initial estimate. So called *self-starting* algorithm, perhaps based on fitting a related/simpler model?

For instance, think of weight-loss ex. Easy to automate initial estimation scheme for initial weight, total loss, ‘half-life’.

## **Nitty-gritty implementation issues**

nls() function in R / S. Lots of flexibility.

Besides starting value issue, need to decide how many derivatives to use: **zero, one, or two.**

In the case of zero, the algorithm will approximate first derivatives numerically. *This is not as easy/reliable as it sounds.*

In the case of one/two derivatives: user-supplied code to evaluate these derivatives or *symbolic manipulation* to do the work [deriv() function]? Latter sounds appealing...

## Confidence intervals

General feeling that  $\pm 1.96SE$  intervals are less reliable in nonlinear problems. WHY?

**ASIDE:** SEs come from observed information matrix, i.e., from second derivative of log likelihood, which we may not have, or want to work at to get.

**FIX:** Let  $l_i$  be contribution of  $i$ -th datapoint to log-likelihood. May wish to replace

$$\hat{I} = - \sum_{i=1}^n l_i''(\hat{\theta}),$$

with

$$\hat{I} = \sum_{i=1}^n \left\{ l_i'(\hat{\theta}) \right\} \left\{ l_i'(\hat{\theta}) \right\}^T,$$



**Alternative # 1:** Invert a test.

Say  $\theta = (\theta_1, \theta_2)$ , want confidence region for  $\theta_1$ .

Could test  $H_0 : \theta_1 = \theta^*$  on the basis of

$$2 \left\{ \max_{\theta_1, \theta_2} l(\theta_1, \theta_2) - \max_{\theta_2} l(\theta^*, \theta_2) \right\} \stackrel{approx}{\sim} \chi_{dim(\theta_1)}^2$$

if  $H_0$  is true.

**Thus, form 95% confidence region for  $\theta_1$  as all values of  $\theta_*$  for which  $H_0 : \theta_1 = \theta^*$  is NOT rejected by this test at significance level 0.05.**

In contrast to SE-based (Wald) interval, such an interval

- may not be symmetric about  $\hat{\theta}_1$
- will behave nicely under reparameterization.

**Alternative #2:** Bootstrap.