

QUIZ # 3

Statistics 305

Term 2, 2006-2007

Tuesday, February 27, 2007

Time: 2:00pm – 2:30pm

Student Name (Please print in caps): SOLUTIONS

Student Number: _____

Notes:

- This quiz has 3 problems on the 4 following pages, plus 1 page of statistical tables. Check to ensure that you have a complete paper.
- The amount each part of each question is worth is shown in [] on the left-hand side of the page.
- Where appropriate, record your answers in the blanks provided on the right-hand side of the page.
- Your solutions **must be justified**; show **all the work** and state **all the reasons** leading to your answer for each question in the space provided immediately under the question.
- Clear and complete solutions are essential; little partial credit will be given.
- This is a closed book exam.
- A single one-sided 8.5 x 11 page of notes is allowed.
- Calculators are allowed (but not for symbolic differentiation or integration).
- No devices (including calculators) that can store text or send/receive messages are allowed.

Problem	Total Available	Score
1.	6	
2.	7	
3.	12	
Total	25	

1. Consider $T = T(X_1, X_2, \dots, X_n)$, any estimator of the parameter θ . As discussed in class, the mean squared error and bias of T are defined as $MSE(T) = E[(T - \theta)^2]$ and $Bias(T) = E(T) - \theta$.

- [4] a) Show that $MSE(T) = Var(T) + [Bias(T)]^2$.

$$\begin{aligned}
 MSE(T) &= E[(T - \theta)^2] \\
 &= E[(T - E(T) + E(T) - \theta)^2] \\
 &= E[(T - E(T))^2 + 2(T - E(T))(E(T) - \theta) + (E(T) - \theta)^2] \\
 &= E[(T - E(T))^2] + 2(E(T) - \theta)E(T - E(T)) + E[(E(T) - \theta)^2] \\
 &= Var(T) + 2(E(T) - \theta) \cdot 0 + (E(T) - \theta)^2 \quad \leftarrow \text{(since there is nothing random in that term)} \\
 &= Var(T) + Bias^2(T) \quad \leftarrow E(\text{constant}) = \text{constant}
 \end{aligned}$$

- [2] b) Explain the practical importance of both terms in the above expression for $MSE(T)$ when evaluating the performance of the estimator T .

- $Var(T)$ describes the precision of the estimator T : how variable is T in repeated samples of size n ? If highly variable, the estimator could take a value far from the target in any particular sample of size n , even if the estimator is unbiased.
- $Bias(T)$ describes the accuracy of the estimator T : how far "off the target" is the estimator, on average, in repeated samples of size n ?

2. Suppose Y has a binomial distribution with parameters n and θ ; that is,

$$P(Y = y) = \binom{n}{y} \theta^y (1 - \theta)^{n-y} \quad \text{for } y = 0, 1, \dots, n.$$

Assume the value of n is known but θ is an unknown parameter.

[5] a) Find the expression for $\hat{\theta}_{ML}$, the maximum likelihood estimator (MLE) of θ based on the single binomial random variable Y .

$$\hat{\theta}_{ML} = \frac{Y}{n}$$

$$L(\theta) = \binom{n}{y} \theta^y (1 - \theta)^{n-y}$$

$$\Rightarrow \ell(\theta) = \log L(\theta) = \log \binom{n}{y} + y \log \theta + (n-y) \log(1 - \theta)$$

$$\Rightarrow \ell'(\theta) = \frac{y}{\theta} - \frac{(n-y)}{1-\theta}$$

Inequality remains the same because $0 < \theta < 1$

$$\ell'(\theta) > 0 \Leftrightarrow \frac{y}{\theta} > \frac{n-y}{1-\theta} \Leftrightarrow y(1-\theta) > (n-y)\theta$$

$$\Leftrightarrow y - y\theta > n\theta - y\theta$$

$$\Leftrightarrow \theta < \frac{y}{n}$$

$$\Rightarrow \hat{\theta}_{ML} = \frac{Y}{n}$$

and it is clear from this that this does correspond to a max of $\ell(\theta)$
 $\Rightarrow$ a max of $L(\theta)$

Alternately, $\ell''(\theta) = -\frac{y}{\theta^2} - \frac{(n-y)}{(1-\theta)^2} < 0$

[2] b) Provide a rough sketch of the likelihood function for $n = 10$ and $y = 5$.

$$L(\theta) = \binom{10}{5} [\theta(1-\theta)]^5 \quad \text{for } 0 < \theta < 1$$

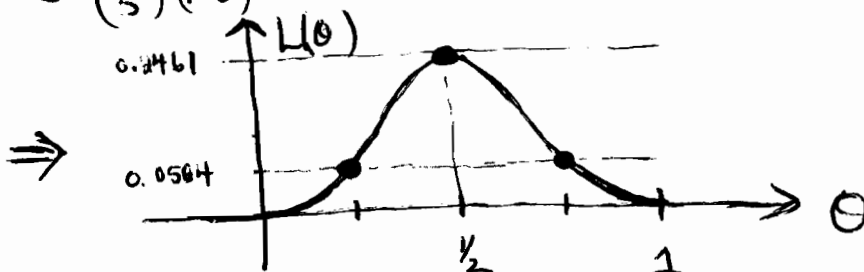
$\Rightarrow$ symmetric around $\theta = \frac{1}{2}$

$$\sim \binom{10}{5} \theta^5 \quad \text{as } \theta \downarrow 0$$

$$\sim \binom{10}{5} (1-\theta)^5 \quad \text{as } \theta \uparrow 1$$

In particular, $\ell''(\frac{y}{n}) = \frac{-n^3}{n(n-y)} < 0$

$\Rightarrow \theta = \frac{y}{n}$ corresponds to a value of θ that maximizes $\ell(\theta) \Rightarrow$ maximizes $L(\theta)$



3. Suppose $X_1, X_2, \dots, X_n$ is a simple random sample from an exponential distribution with a rate of θ ; that is, from the density function:

$$f_{\theta}(x) = \theta \exp(-\theta x) \quad \text{for } x > 0.$$

- [3] a) Find the expression for $\hat{\theta}_{MM}$, the method of moments estimator (MME) of θ .

$$\hat{\theta}_{MM} = \frac{1}{\bar{X}}$$

This density is a $\text{Gamma}(1, \theta) \Rightarrow E(X) = \frac{1}{\theta}$, the 1st population moment
 μ_1

$$\Rightarrow \theta = \frac{1}{\mu_1}$$

$$\Rightarrow \hat{\theta}_{MM} = \frac{1}{\hat{\mu}_1} \leftarrow (\text{the 1st sample moment})$$

$$\Rightarrow \hat{\theta}_{MM} = \frac{1}{\bar{X}} \quad \text{Var}(\hat{\theta}_{MM}) \approx \frac{\theta^2}{n}$$

- [5] b) Find an approximate expression for the variance of the MME $\hat{\theta}_{MM}$.

In any random variable Y , the delta method yields

$$\text{Var}[g(Y)] \approx [g'(E(Y))]^2 \cdot \text{Var}(Y)$$

We have $\hat{\theta}_{MM} = \frac{1}{\bar{X}} \equiv g(\bar{X})$, where $g(x) = \frac{1}{x}$

$$\Rightarrow \text{Var}(\hat{\theta}_{MM}) \approx \left[-\frac{1}{E(\bar{X})^2} \right]^2 \cdot \text{Var}(\bar{X}) \quad \rightarrow g'(x) = -\frac{1}{x^2}$$

$$= \left[\frac{1}{E(X)} \right]^4 \cdot \frac{\text{Var}(X)}{n}$$

But $\text{Var}(\bar{X}) \stackrel{\text{Always!}}{=} \frac{1}{n} \text{Var}(X)$

and $\text{Var}(X) = \frac{1}{\theta^2}$ for $\mathcal{E}(1, \theta)$

$$= \theta^4 \cdot \frac{1}{\theta^2 n} = \frac{\theta^2}{n}$$

Also $E(\bar{X}) \stackrel{\text{Always!}}{=} E(X) = \frac{1}{\theta}$ for $\mathcal{E}(1, \theta)$

[4] c) Find a second-order approximate expression for the bias of the MME $\hat{\theta}_{MM}$. $\text{Bias}(\hat{\theta}_{MM}) \approx \frac{0}{n}$

In any random variable Y , the delta method yields

As in b), $E[g(Y)] \approx g(E(Y)) + \frac{1}{2} g''(E(Y)) \cdot \text{Var}(Y)$

We have $\hat{\theta}_{MM} \equiv g(\bar{X})$, where $g(x) = \frac{1}{x} \Rightarrow g'(x) = -\frac{1}{x^2}$

$$\Rightarrow E(\hat{\theta}_{MM}) \approx \frac{1}{E(\bar{X})} + \frac{1}{2} \frac{2}{[E(\bar{X})]^3} \text{Var}(\bar{X})$$

$$= 0 + 0^3 \cdot \frac{1}{0^2 \cdot n}$$

$$= 0 + \frac{0}{n}$$

$$\Rightarrow \text{Bias}(\hat{\theta}_{MM}) = E(\hat{\theta}_{MM}) - \theta$$

$$\approx 0 + \frac{0}{n} - 0 = \underline{\underline{\frac{0}{n}}}$$

Note that $\text{Bias}^2(\hat{\theta}_{MM}) \approx \frac{0^2}{n^2}$

In particular, $\text{Bias} \rightarrow 0$ as $n \rightarrow \infty$
 $\rightarrow$ asymptotically unbiased!

$$\begin{aligned} \Rightarrow \text{MSE}(\hat{\theta}_{MM}) &= \text{Var}(\hat{\theta}_{MM}) + \text{Bias}^2(\hat{\theta}_{MM}) \\ &\approx \frac{0^2}{n} + \frac{0^2}{n^2} = \frac{0^2}{n} \text{ to } \end{aligned}$$

$\Rightarrow$ In large n , $\text{MSE}(\hat{\theta}_{MM})$ is completely determined by the variance because $\hat{\theta}_{MM}$ has bias $\sim 1/n$ leading order (for large n)

THE END

$\uparrow$ (This is typical for reasonable estimators: if they are biased, the bias contribution to MSE is negligible for large n .)