

Module 3: Gaussian Process Parameter Estimation, Prediction Uncertainty, and Diagnostics

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Adapted from materials prepared by Jerry Sacks and Will Welch for various short courses

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Outline of Topics

- 1 Estimating the Parameters of the GP Model
- 2 Case Study: G-Protein Computer Experiment
- 3 Measuring Prediction Accuracy
- 4 GP Diagnostics
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Parameters of the Gaussian Process (GP) Model

Recall from Module 2 that the Gaussian process prior for $y(\mathbf{x}) = y(x_1, \dots, x_d)$ has hyper-parameters:

- **mean**, μ ,
- **variance**, σ^2
- **correlation parameters**, e.g., $\theta_1, \dots, \theta_d$ and p_1, \dots, p_d for the power-exponential correlation function,

$$R(\mathbf{x}, \mathbf{x}') = \prod_{j=1}^d \exp(-\theta_j |x_j - x'_j|^{p_j}).$$

- Their values will be chosen to be consistent with the computer-model runs.



Maximum Likelihood

- Recall also that $y(\mathbf{x})$ is assumed to be Gaussian.
- Hence, $\mathbf{y} = [y(\mathbf{x}^{(1)}), \dots, y(\mathbf{x}^{(n)})]^T$, the data from the computer model, are a sample from a multivariate-normal distribution.
- The **likelihood**, $L(\mathbf{y} \mid \mu, \sigma^2, \theta_1, \dots, \theta_d, \rho_1, \dots, \rho_d)$, is

$$\frac{1}{(2\pi\sigma^2)^{n/2} \det^{1/2}(\mathbf{R})} \exp\left(-\frac{1}{2\sigma^2}(\mathbf{y} - \mu\mathbf{1})^T \mathbf{R}^{-1}(\mathbf{y} - \mu\mathbf{1})\right).$$

- **Maximum likelihood estimation (MLE)** chooses the hyper-parameters to maximize this.
- Or use **Bayes' rule** to get a **posterior distribution** for the hyper-parameters and for predictions of $y(\mathbf{x})$ (see Appendix A).



Maximum Likelihood: Computation

For fixed correlation parameters,

$$\hat{\mu} = \frac{\mathbf{1}^T \mathbf{R}^{-1} \mathbf{y}}{\mathbf{1}^T \mathbf{R}^{-1} \mathbf{1}}$$

and

$$\hat{\sigma}^2 = \frac{1}{n} (\mathbf{y} - \hat{\mu} \mathbf{1})^T \mathbf{R}^{-1} (\mathbf{y} - \hat{\mu} \mathbf{1})$$

The likelihood function (with $\hat{\mu}$ and $\hat{\sigma}^2$ substituted) has to be **numerically maximized w.r.t. the correlation parameters.**



G-Protein Computer Model

Biosystems model for so-termed ligand activation of G-protein in yeast.

$d = 4$ input variables

- x is concentration of ligand
- u_1, \dots, u_8 is a vector of 8 kinetic parameters (only u_1 , u_6 , and u_7 are varied)

Output variable

- y is the normalized concentration of part of the complex



G-Protein System Dynamics: Differential Equations

- 1 $\dot{\eta}_1 = -u_1\eta_1 + u_2\eta_2 - u_3\eta_1 + u_5$
- 2 $\dot{\eta}_2 = u_1\eta_1 - u_2\eta_2 - u_4\eta_2$
- 3 $\dot{\eta}_3 = -u_6\eta_2\eta_3 + u_8(G_{\text{tot}} - \eta_3 - \eta_4)(G_{\text{tot}} - \eta_3)$
- 4 $\dot{\eta}_4 = u_6\eta_2\eta_3 - u_7\eta_4$
- 5 $y = (G_{\text{tot}} - \eta_3)/G_{\text{tot}}$

where

- η_1, \dots, η_4 are concentrations of 4 chemical species and $\dot{\eta}_1 \equiv \frac{\partial \eta_1}{\partial t}$, etc.
- $G_{\text{tot}} = (\text{fixed})$ total concentration of G-protein complex after 30 seconds



Inputs and Code Runs

Input variables

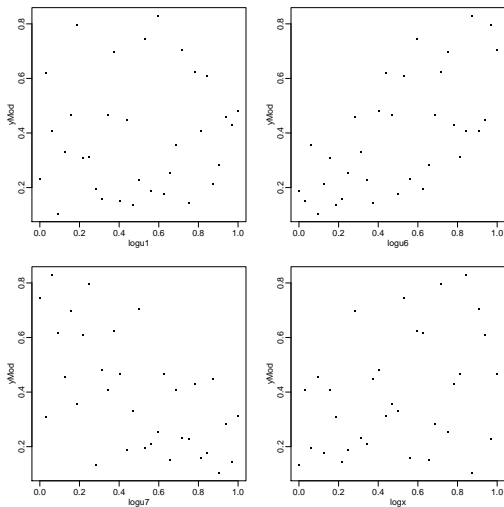
- $d = 4$ variables
- Work with $\log(x)$, $\log(u_1)$, $\log(u_6)$, $\log(u_7)$.
- i.e., what we called the \mathbf{x} vector before is $\log(x)$, $\log(u_1)$, $\log(u_6)$, and $\log(u_7)$ here
- All input variable ranges are normalized to $[0, 1]$ on the log scale

Number of runs

- $n = 33$
(this choice and the design for the 33 runs is described in Module 4)



Computer Model Data



Gaussian Process (GP) Model

$y(\mathbf{x})$ is a realization of a Gaussian process with:

- mean μ
- variance σ^2
- correlations given by

$$\text{Cor}(y(\mathbf{x}), y(\mathbf{x}')) \equiv R(\mathbf{x}, \mathbf{x}') = \prod_{j=1}^4 e^{-\theta_j |x_j - x'_j|^{p_j}}.$$

The parameters in red need to be estimated.



Maximum Likelihood Estimates

- $\hat{\mu} = 0.36$
- $\hat{\sigma}^2 = 0.51$

Variable	$\hat{\theta}$	$\hat{\rho}$
$\log(x)$	0.929	1.98
• $\log(u_1)$	0.179	2
$\log(u_6)$	0.082	2
$\log(u_7)$	0.083	2

- It is difficult to interpret the magnitudes of the estimates. (we will revisit this example in Module 5 and do a sensitivity analysis).



“Plug-In” Prediction and Standard Error

Replace all hyper-parameters by their MLEs in the conditional mean and variance formulas:

$$\text{prediction of } y(\mathbf{x}) = \hat{y} = \hat{\mathbf{m}}(\mathbf{x}) = \hat{\mu} + \mathbf{r}^T(\mathbf{x})\mathbf{R}^{-1}(\mathbf{y} - \hat{\mu}\mathbf{1}).$$

and

$$\text{estimated variance of prediction} = \hat{\mathbf{v}}(\mathbf{x}) = \widehat{\sigma}^2(1 - \mathbf{r}^T(\mathbf{x})\mathbf{R}^{-1}\mathbf{r}(\mathbf{x})).$$

(\mathbf{R} and $\mathbf{r}(\mathbf{x})$ are also estimates.)

The plug-in estimated variance **ignores uncertainty in estimating the hyper-parameters**. It can be adapted to include **uncertainty from estimating μ** :

$$\hat{\mathbf{v}}(\mathbf{x}) = \widehat{\sigma}^2 \left(1 - \mathbf{r}^T(\mathbf{x})\mathbf{R}^{-1}\mathbf{r}(\mathbf{x}) + \frac{[1 - \mathbf{1}^T\mathbf{R}^{-1}\mathbf{r}(\mathbf{x})]^2}{\mathbf{1}^T\mathbf{R}^{-1}\mathbf{1}} \right).$$

This plug-in formula is often used to give a **standard error**, i.e.,
 $s(\mathbf{x}) = \sqrt{\hat{\mathbf{v}}(\mathbf{x})}$.



Measures of Accuracy

- We could rely on the **standard error**, $\sqrt{\hat{v}(\mathbf{x})}$.
- If we have m **test data** observations, the root mean squared error (RMSE) of prediction is

$$\text{RMSE} = \sqrt{\frac{1}{m} \sum_{\text{test pts}} (\hat{y} - y(\mathbf{x}))^2}.$$

But rarely available.

- **Cross validation (CV)**



Cross Validation (CV)

Let $\mathbf{x}^{(i)}$ denote \mathbf{x} for run i in the data ($i = 1, \dots, n$). For run i :

- The **cross validated prediction** of $y(\mathbf{x}^{(i)})$ is

$$\hat{y}_{-i}(\mathbf{x}^{(i)}),$$

i.e., $\hat{y}(\mathbf{x}) = \hat{m}(\mathbf{x})$ computed from the $n - 1$ runs **excluding run i** .

- The **cross validated standard error** of $\hat{y}_{-i}(\mathbf{x}^{(i)})$ is

$$s_{-i}(\mathbf{x}^{(i)}),$$

i.e., $s(\mathbf{x}) = \sqrt{\hat{v}(\mathbf{x})}$ computed from the $n - 1$ runs excluding run i .

- The **cross-validated residual** for run i is

$$y(\mathbf{x}^{(i)}) - \hat{y}_{-i}(\mathbf{x}^{(i)}).$$

- The standardized cross-validated residual for run i is

$$\frac{y(\mathbf{x}^{(i)}) - \hat{y}_{-i}(\mathbf{x}^{(i)})}{s_{-i}(\mathbf{x}^{(i)})}.$$

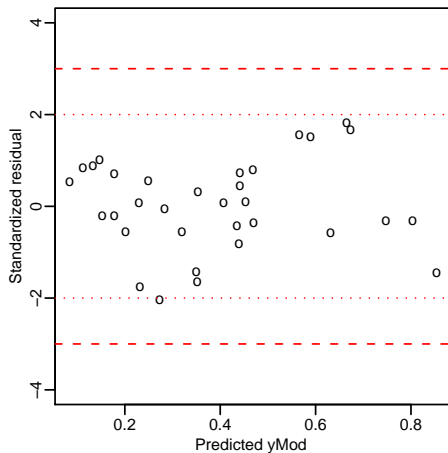
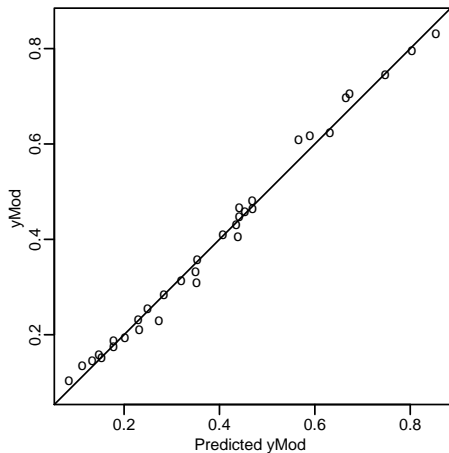


Diagnostic Plots

- Plot the cross-validated residuals to assess the overall **magnitude of error**.
- Plot the standardized cross-validated residuals to assess the **validity of the standard error** for individual predictions.



G-Protein Diagnostic Plots



Cross Validation: Numerical Summaries

Magnitude of error

- The cross-validated root mean squared error is

$$CVRMSE \equiv \sqrt{\frac{1}{n} \sum (y(\mathbf{x}^{(i)}) - \hat{y}_{-i}(\mathbf{x}^{(i)}))^2} = .020.$$

- Maximum cross-validated residual is .044
- Fairly accurate relative to a range of about 0.7 in y

Standard errors?

- $\frac{y(\mathbf{x}^{(i)}) - \hat{y}_{-i}(\mathbf{x}^{(i)})}{s_{-i}(\mathbf{x}^{(i)})}$ for $i = 1, \dots, n$ are roughly in $(-2, 2)$
- Standard errors look reliable.



Fast and Slow CV

- When run i is removed, the hyper-parameters should be re-estimated.
- For computational reasons the correlation parameters are often **not** updated (it is cheap to update the estimates of μ and σ^2), producing a “fast” CV.
- For “slow” CV, do say 10-fold cross-validation, re-estimating all hyper-parameters.
- The agreement between “fast” CVRMSE and “slow” CVRMSE is often good.
- The agreement between “fast” CVRMSE and the RMSE from test points has been good in examples.



Module Summary

- The GP model has to be “tuned” to data so that its properties match those of the computer model.
- Tuning (fitting) the GP by maximum likelihood is computationally feasible for up to about $n = 1000$ runs and $d = 50$ input variables.
- GP model gives an approximation and a measure of accuracy.
- The measure of accuracy (standard error) can be checked for validity by cross validation.



Appendix A: Bayesian Treatment of the Hyper-parameters

- Posterior distribution of the hyper-parameters (“hyper” below), μ , σ^2 , $\theta_1, \dots, \theta_d$, etc., of the GP
 - From Bayes rule, given the data \mathbf{y}

$$p(\text{hyper} | \mathbf{y}) \propto \pi(\text{hyper})L(\mathbf{y} | \text{hyper}),$$

- $\pi(\text{hyper})$ is the prior for hyper
- $L(\mathbf{y} | \text{hyper})$ is the multivariate normal likelihood.
- Predictive distribution for $y(\mathbf{x})$ at a “new” \mathbf{x}
 - $p(y(\mathbf{x}) | \mathbf{y}) = \int p(y(\mathbf{x}) | \mathbf{y}, \text{hyper})p(\text{hyper} | \mathbf{y}) d\text{hyper}$
 - Usually, the integration is not carried out explicitly.
 - Rather, properties such as the posterior predictive mean and variance of $p(y(\mathbf{x}_0) | \mathbf{y})$ are obtained by MCMC sampling of the posterior distribution for the hyper-parameters, $p(\text{hyper} | \mathbf{y})$.

